

# Monte Carlo and nonlinearities

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## Abstract

The Monte Carlo method is widely used to numerically predict systems behaviour. However, its powerful incremental design assumes a strong premise which has severely limited application so far: the estimation process must combine linearly over dimensions. Here we show that this premise can be alleviated by projecting nonlinearities on a polynomial basis and increasing the configuration-space dimension. Considering phytoplankton growth in light-limited environments, radiative transfer in planetary atmospheres, electromagnetic scattering by particles and concentrated-solar-power-plant productions, we prove the real world usability of this advance on four test-cases that were so far regarded as impracticable by Monte Carlo approaches. We also illustrate an outstanding feature of our method when applied to sharp problems with interacting particles: handling rare events is now straightforward. Overall, our extension preserves the features that made the method popular: addressing nonlinearities does not compromise on model-refinement nor system complexity and convergence rates remain independent of dimension.

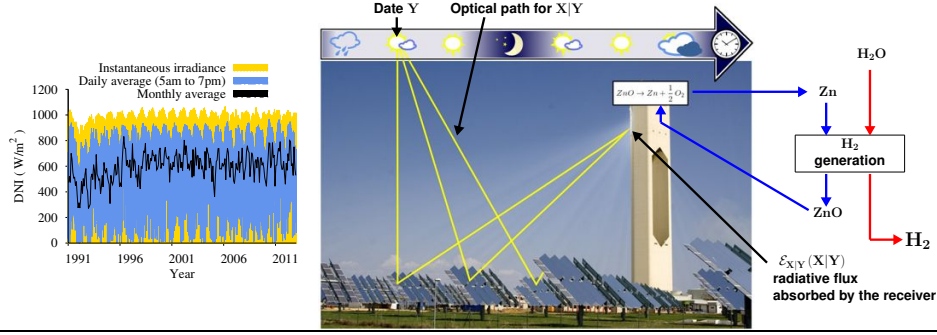
The standard Monte Carlo (MC) method is a method to predict a physical observable by numerically estimating a statistical expectation over a multi-dimensional configuration space[1]. The reason why this method is so popular in all fields of scientific research is its intuitive nature. In the most current practice, simulation tools are designed in direct relation with the pictures available in each discipline and later refinements are gradual and straightforward. Model refinements merely extend sampling to new suitable dimensions. The method is nonetheless mathematically rigorous: specialists specify observables that are implicitly translated into integral quantities that are estimated using random sampling in each direction of the configuration space. This choice of a statistical viewpoint is highly powerful because the algorithm can be designed directly from the description of the system, whether it is deterministic or not, with no reworking nor approximation.

Let us illustrate how MC is used in engineering with a typical example: the optimal design of a concentrated solar plant[2] (see Fig.1-a). The power collected by the central-receiver results from all the rays of sunlight that reach it after reflection by heliostats, so it depends on the complex geometry of the heliostats. Moreover, the heliostats being adaptively oriented to follow the sun position, they can mask one another at some times of the day. To estimate by MC the received power at one time of interest, i.e. for a given geometry of the heliostats: pick an optical path among the ones that join the sun to the central-receiver via an heliostat, check whether this path is shadowed or blocked by another heliostat, and retain a Monte Carlo weight equal to 0 or 1 depending on the transmission success. Let  $\mathbf{X}$  be the random variable denoting the transmission success. The collected fraction of the available sun power is then the expectation  $\mathcal{E}_{\mathbf{X}}(\mathbf{X})$  of  $\mathbf{X}$  and can be evaluated with no bias as the average of such weights over a large number of sampled paths.

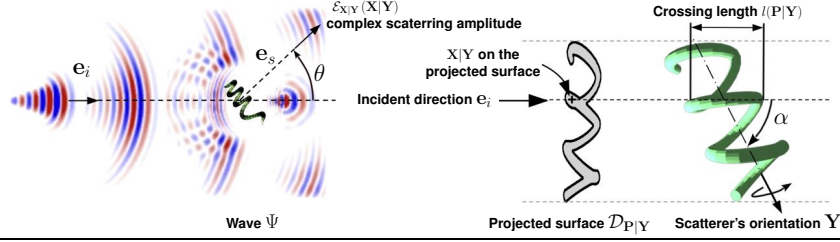
This approach robustly complies with expanded descriptions of the physical observable to be addressed. For instance, the fraction of the available sun power collected on average over the entire lifetime of the plant (operating typically 30 years) can be predicted as the expectation over time of  $\mathcal{E}_{\mathbf{X}}(\mathbf{X})$  varying in time. Denoting  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  the collected fraction at random time  $\mathbf{Y}$  within the 30 years, the time-averaged fraction is given by  $\mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})) = \mathcal{E}_{\mathbf{Y},\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$ . The basic algorithm above can then be encapsulated within time sampling: first pick a date for  $\mathbf{Y}$ , then pick a path at that date for  $\mathbf{X}|\mathbf{Y}$ . In the end, estimate  $\mathcal{E}_{\mathbf{Y},\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  by computing the average of the transmission success over all combined pairs (date, path). Meanwhile, one can as well account for sun power fluctuations by estimating the atmospheric transmission at each picked date. The choice of the statistical viewpoint thus allows to incorporate into one single statistical question as many ingredients as it takes: the geometrical complexity of the heliostats[3], the daily course of the sun, the seasonal scale as well as the hour scale weather fluctuations[6]. Remarkably, the latter question is nearly as simple to address as the estimation of the power collected at one single date: the algorithmic design can map the full conceptual description, yet computational costs are poorly affected. Contrastingly, deterministic approaches would translate into unpractical computation times or require simplified and approximate descriptions, so MC has become the only practical solution in many such engineering contexts. Becoming such a standard practice, MC has prompted numerous theoretical developments[6, 7, 5, 8].

Figure 1 — Complex systems with non linear outputs: four real-world examples.

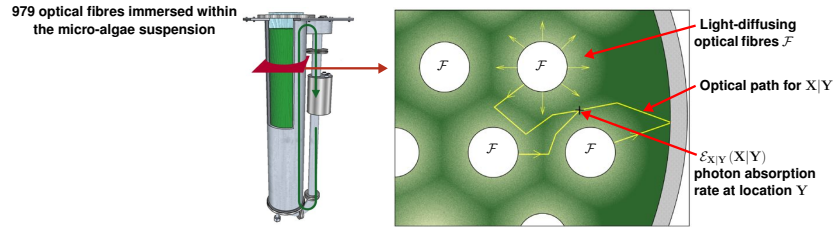
**a Solar Plant:**  $C(p) = \mathcal{E}_Y \left( \exp \left( C_4 \left( \mathcal{E}_{X|Y}(X|Y) \right)^{-\alpha} \right) \sum_{m=0}^3 C_m \left( \mathcal{E}_{X|Y}(X|Y) \right)^{(m\alpha-1)} \right)$



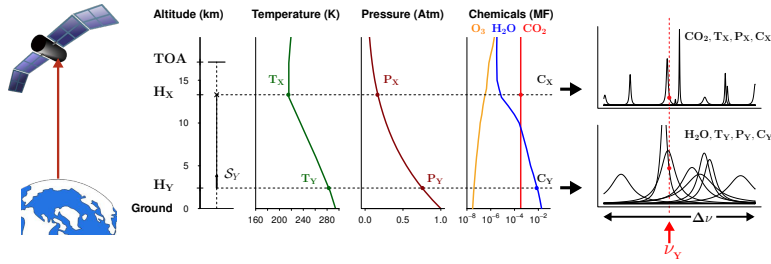
**b Wave scattering:**  $S(e_s) = \mathcal{E}_Y \left( \left| \mathcal{E}_{X|Y}(X|Y) \right|^2 \right)$



**c Biomass production:**  $R(C) = \mathcal{E}_Y \left( C \left( \alpha \frac{K \mathcal{E}_{X|Y}(X|Y)}{K + \mathcal{E}_{X|Y}(X|Y)} - \beta \frac{K_r}{K_r + \mathcal{E}_{X|Y}(X|Y)} \right) \right)$



**d Atmospheric radiative transfert:**  $I(\Delta\nu) = \mathcal{E}_Y \left( S_Y \exp(-\mathcal{E}_{X|Y}(X|Y)) \right)$



**Fig. 1 Complex systems with nonlinear outputs: four real world examples.**

**a, Solar-driven high-temperature thermal reduction of zinc oxide, as the first part of a two step water splitting cycle.** Photons emitted from the sun are reflected on heliostats and concentrated at the entrance of the chemical reactor in which  $ZnO$  dissociation is carried out. Depending on their transmission success  $\mathbf{X}$ , the solar power  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  absorbed by the receiver at a random instant  $\mathbf{Y}$  of lifetime determines the nonlinear chemical conversion rate of the reaction  $ZnO \rightarrow Zn + \frac{1}{2}O$ . Here we address the estimation of the annual solar-plant's conversion rate  $C(\mathbf{p})$  at different earth locations  $\mathbf{p}$ , by averaging the instantaneous conversion rates over the statistics of sun position and incident Direct Normal Irradiance (DNI), which fluctuates with time and weather (see also EDF1 and SI1).

**b, Wave scattering by a complex-shaped and optically-soft scatterer (cyanobacterium *Arthrospira*).** An incident plane wave with propagation direction  $\mathbf{e}_i$  is scattered by the helical cyanobacterium. The bacterium has low relative refractive index and is much larger than wavelength (optically-soft particle). The complex scattering-amplitude  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  in the forward directions is the sum of secondary waves contributions  $\mathbf{X}|\mathbf{Y}$  (interference) originating from the scatterer projected surface. This surface depends on the scatterer orientation  $\mathbf{Y}$ . Here we address the estimation of  $S(\mathbf{e}_s)$  the single-scattering differential cross-section in direction  $\mathbf{e}_s$  for a suspension of particles, assuming independent scattering, by averaging the squared modulus of  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  over the statistics of orientations  $\mathbf{Y}$  (see also EDF2 and SI2).

**c, Phytoplankton growth in light-limited environments.** Phytoplankton is put to grow in a continuous stirred tank photobioreactor internally illuminated by optical fibres  $\mathcal{F}$  immersed in the culture. The local rate of photon absorption  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  at location  $\mathbf{Y}$  is the average of the contributions  $\mathbf{X}|\mathbf{Y}$  of every optical path from fibres to  $\mathbf{Y}$  through the scattering and absorbing suspension.  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  determines the nonlinear photosynthetic growth-rate at location  $\mathbf{Y}$ . Here, we address the Monte Carlo estimation of  $R(C)$  the global growth-rate in the whole culture volume as a function of biomass concentration  $C$ , by averaging the local rate over locations in the volume (see also EDF3 and SI3).

**d, Atmospheric radiative transfer : top-of-atmosphere (TOA) specific intensity (from earth toward the outer space).** Photons emitted by the atmosphere at infrared frequencies are due to random emission-transitions  $\mathbf{Y}$ , from an upper to a lower energy-state, of mainly  $CO_2$  and  $H_2O$  molecules of concentration  $C_{\mathbf{Y}}$  at altitude  $H_{\mathbf{Y}}$ . The corresponding source  $\mathcal{S}_{\mathbf{Y}}$  depends on the thermodynamic state of the atmosphere, mainly temperature  $T_{\mathbf{Y}}$  (defining the energy-state population) and pressure  $P_{\mathbf{Y}}$  (defining most of the line-width, i.e. the uncertainty of the emission-frequency  $\nu_{\mathbf{Y}}$ ). This source is then exponentially extinct by atmospheric absorption, i.e. by all random absorption-transitions  $\mathbf{X}|\mathbf{Y}$ , from a lower to an upper energy-state, occurring at altitude  $H_{\mathbf{X}|\mathbf{Y}}$  between  $H_{\mathbf{Y}}$  and the top of atmosphere (see also EDF4 and SI4).

Still, MC cannot handle *any* question to date. As a matter of fact, it has been identified early on that “the extension of Monte Carlo methods to nonlinear processes may be impossible”[9] and it is a prevalent opinion nowadays that “Monte Carlo methods are not generally effective for nonlinear problems mainly because expectations are linear in character”[10] so that “a nonlinear problem must usually be linearized in order to use Monte Carlo technique”[11]. We are aware of only one attempt so far to bypass this failing: the proposition by Vajargah[12] or Dimov[1] of using branching processes to solve Fredholm-type integral equations with polynomial nonlinearity.

Unluckily, most real world problems are nonlinear. And indeed, if the question were now to evaluate the final return on investment of the solar plant, namely how much electrical power it would deliver over its lifetime, standard MC would fail because the instantaneous conversion efficiency from collected solar-power to electrical power is not linear. Let us consider, as a toy example, a basic nonlinear case where the electrical power would be proportional to the square of the instantaneous collected solar power  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  at date  $\mathbf{Y}$ . In Monte-Carlo terms, the question would be then to estimate  $\mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})^2)$  over the plant lifetime. In this case, the optical and temporal expectations cannot be combined any more because it would be wrong to first estimate, as above, the total solar power collected over lifetime, and then apply the conversion efficiency in the end (basically,  $\mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})^2) \neq \mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y}))^2$  the same way as  $a^2 + b^2 \neq (a + b)^2$ ). Instead, one would have to sample dates (say  $M$  dates, millions over 30 years), estimate the solar power collected at each date by averaging over numerous optical paths transmission successes (say  $N$  paths, millions for each date), apply the nonlinear conversion to the result at that date, and then average over all dates. Doing so, MC would now require  $M \times N$  samples, and, even worse, further levels of complexity (each adding a nonlinearity to the problem) would similarly multiply the computation time. Moreover, the result would be biased due to the finite sampling sizes of the innermost dimensions. In short, MC distinctive features are dismissed and exact life-time integration looks impossible.

Bearing in mind our earlier theoretical works about MC integral formulations[2], we found a way to bypass this obstacle for a large class of nonlinear problems, based on the very statistical nature of MC. In the case of our toy example, we use the fact that:

$$\mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})^2) = \mathcal{E}_{\mathbf{Y},(\mathbf{X}_1, \mathbf{X}_2)|\mathbf{Y}}(\mathbf{X}_1 \mathbf{X}_2|\mathbf{Y}) \quad (1)$$

where  $\mathbf{X}_1$  and  $\mathbf{X}_2$  are two independent variables identically distributed as  $\mathbf{X}$  (see Methods). Translated in sampling algorithm, the solution is then to sample optical paths in pairs  $(\mathbf{X}_1, \mathbf{X}_2)|\mathbf{Y}$  (instead of millions) at each sampled date and now retaining the pair product  $\mathbf{X}_1\mathbf{X}_2|\mathbf{Y}$  of their transmission successes. The optical and temporal statistics can then actually be sampled together and yield the unbiased result with no combinatorial explosion. This reformulation can be generalised to any nonlinearity of polynomial shape. First, monomials of any degree can indeed be estimated using the same sampling property as the one used above for  $n = 2$ :

$$\mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})^n) = \mathcal{E}_{\mathbf{Y},(\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n)|\mathbf{Y}}(\mathbf{X}_1 \mathbf{X}_2 \dots \mathbf{X}_n|\mathbf{Y}) \quad (2)$$

where the  $\mathbf{X}_i$  are  $n$  independent random variables identically distributed as  $\mathbf{X}$ . For any monomial of degree  $n$ , the expectation can then be computed by sampling series of  $n$  independent realisations of  $\mathbf{X}|\mathbf{Y}$ , and averaging the series products. The linear case, solved by standard MC, corresponds to  $n = 1$ . Secondly, since polynomials are just linear combinations of monomials, the expectation for any polynomial function of  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  of degree  $n$  can be translated into a Monte Carlo algorithm,

sampling first a degree in the polynomial, and then sampling as many independent realisations of  $\mathbf{X}|\mathbf{Y}$  as this random degree (see *Methods*). For a polynomial function of degree  $n$ , the corresponding Non-Linear Monte Carlo (NLMC) algorithm is then:

- pick a sample  $y$  of  $\mathbf{Y}$ ,
- pick a monomial degree value  $d \leq n$ ,
- draw  $d$  independent samples of  $\mathbf{X}|\mathbf{Y} = \mathbf{y}$  and retain their product,

repeat this sampling procedure and compute the estimate as the average of retained products.

Moreover, if polynomial forms of any dimension are now solvable with no approximation, so is the projection of any nonlinear function on a polynomial basis of any dimension, even of infinite dimension if needed (full details using the Taylor expansion are given in *Methods*). As a result, any hierarchy of nested statistical processes that combine nonlinearly can now be, in theory, exactly addressed within the Monte Carlo framework. The deep rationale of our reformulation of nonlinear expectations can be grounded on the fundamentals of functional statistics: nonlinear processes are formally equivalent to infinite-dimension processes and MC algorithms are well known to smartly address, by design, expectations over domains of infinite dimension.

To our knowledge, this analysis has never been made. It has yet major practical consequences for real world problems, provided the polynomial sampling, which is the price to be paid for tackling nonlinearities exactly, remains tractable. For instance, let's go back to our illustrative solar-power-plant example, and use now the actual expression for the conversion rate and its Taylor expansion: for each date, once a sun-position and climate conditions have been fixed, one would have to pick a random number of independent optical paths (instead of one optical path in the linear case), keep the product of their transmission success, and make the average of many such products in the end. Doing so, it becomes possible to integrate hourly solar input fluctuations over 30 years in the full geometry of a kilometre-wide heliostats field in order to optimise the nonlinear solar-to-electric conversion over the plant lifetime (Fig. 1A). The same line of thought can be used to predict wave scattering by a tiny scale complex-shaped scatterer[7] such as an helicoidal cyanobacterium (Fig. 1B). The biomass production example (Fig. 1C), where incoming light favours the photosynthetic growth that in turn blocks the incoming light, illustrates how our method also handles nonlinear feedbacks[14]. Finally, with the estimation of Earth radiative cooling (Fig. 1D), we reproduce quite classical results, yet with a pure statistical approach: by sampling directly the state transitions of greenhouse gases, we avoid costly deterministic computations that the standard linear Monte Carlo approach requires in order to by-pass the nonlinearity of Beer Extinction Law[15]. In each of the four cases, it appeared that the additional computations were well affordable using only ordinary computing power (the complete physical descriptions of the four problems, the nonlinearities involved and their translation in NLMC can be found in their respective Extended Data Figures and Supplemental Information, Solar Plant: EDF1, SI1 ; Complex-shaped Scatterer: EDF2, SI2 ; Biomass production: EDF3, SI3 ; Earth radiative cooling: EDF4, SI4).

From a theoretical point of view, in these four cases, the model is directly enunciated in statistical terms, defining two random variables  $\mathbf{X}$  and  $\mathbf{Y}$  from the start. More broadly, standard MC practice can start as well from a deterministic description (see *Methods*), most commonly from a linear partial differential equation (PDE). The formal equivalence between the solution of a linear PDE and the expectation of a random variable has been established for long[16]. And indeed PDE-to-MC

translations are essential to nanoscale mechanics (Quantum Monte Carlo[17]) or nuclear sciences. NLMC allows such translations for nonlinear PDE.

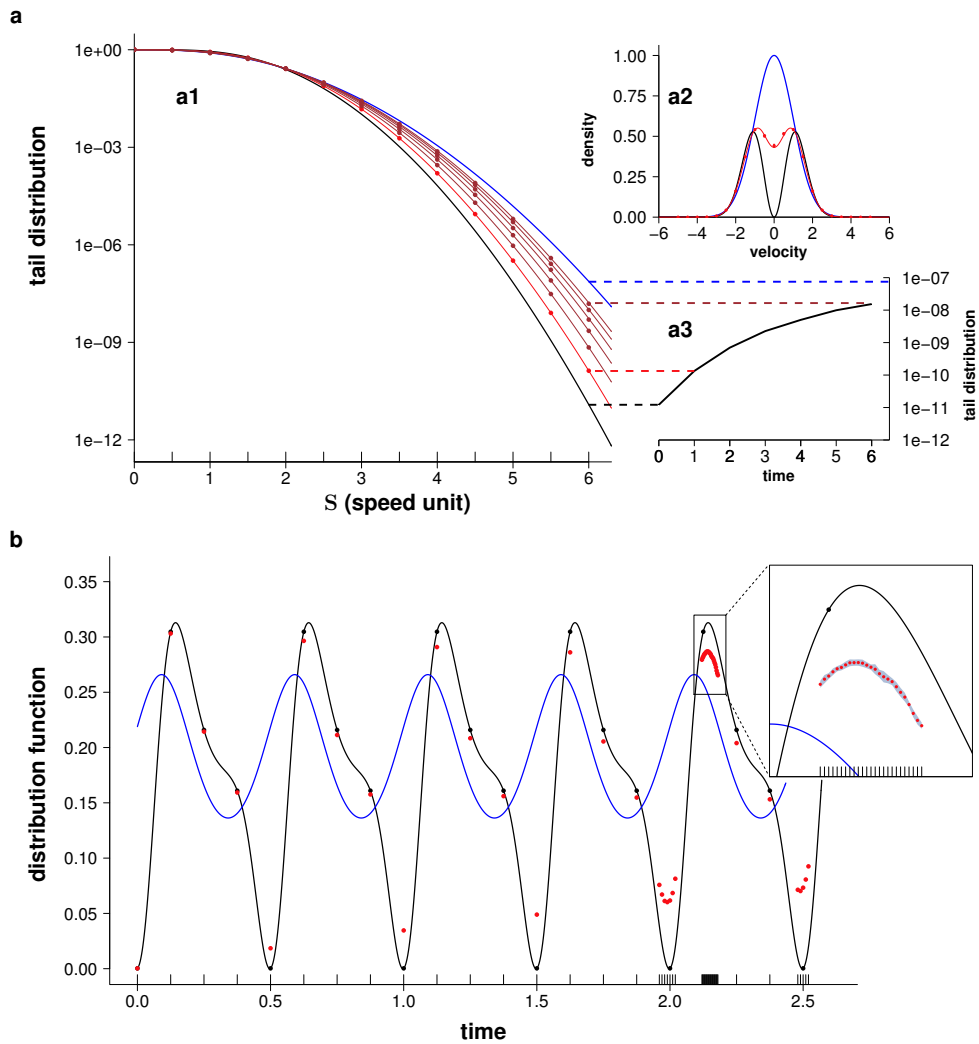
As a touchstone, we address a prominent example of nonlinear PDE in statistical physics, the Boltzmann equation which governs the spatiotemporal density of interacting particles in a dilute gas (full details in SI5). The corresponding physics is easy to visualise: a particle just follows its ballistic flight until it collides with another particle. The collisions are considered as instantaneous and only modify the two particles velocities. The equation for the evolution of particle density in phase-space (position, velocity) is nonlinear because the collision rate depends on the density itself. In order to project this nonlinearity on the proper polynomial basis of infinite dimension, this PDE is first translated into its Fredholm integral counterpart (a step reminiscent of the aforementioned Dimov’s proposition[1]). This Fredholm integral expresses the density in the phase space at some location for some velocity at some time, as if putting a probe in space-time. It is estimated by Monte Carlo, tracking the dynamics backward in time up to the initial condition (or boundary conditions). Importantly, such a probe estimation does not require the exhaustive resolution of the whole field at previous times: like in standard backward MC for linear transport, the information about previous states of the field is reconstructed along each path only where and when it is required[20]. The contrast with linear MC is that nonlinearity due to collisions translate into branching paths.

This extension deals very efficiently with extremely rare events because it preserves an essential feature of MC: avoiding time / space / velocity discretisation[19, 20, 21], very low densities can be estimated with no bias, and the only source of uncertainty is the finite number of sampled events (i.e. the confidence interval around the estimated density). As a test bed, we consider a case for which analytical solutions have been published: the early Krook’s analysis of the distribution of speeds in far out-of-equilibrium conditions[22, 23]. Krook’s analysis was outstanding because it provided an analytical solution to a problem which looked impossible to solve numerically: events with the greatest consequences, namely particles with the highest energies (i.e. high speed particles, of tremendous importance in nuclear chemistry) lie far in the tail of speed distribution and have very low probability of occurrence (rare events). Using our NLMC design, the fractions of particles which have a kinetic energy higher than  $10^6$  times the average value, and which correspond to a fraction as low as  $10^{-11}$  of the total, can now be quantified as accurately as desired, and fit perfectly the analytical solution (fig. 2a).

Being validated on Krook’s case, this extension opens the way to solve systems for which no analytical solutions are available. As an example, we consider now a fully spatialised system in which the particles are confined by an outside harmonic potential, leading to a so-called *breathing mode* of the gas density. Such a solution to the Boltzmann equation had been early identified by Boltzmann himself[24], but has been recently revisited and generalised in the context of shortcut to adiabaticity techniques for classical gases[25]. Exact solutions are available only under the constraint that gas is at local equilibrium, in which case density displays a permanent oscillation. Here again, these analytical solutions are exactly recovered. Yet, NLMC allows to go beyond this constraint and to explore the gas behaviour when the local equilibrium constraint is alleviated: starting from a state far from local equilibrium, it is now possible to estimate how fast the velocities redistribution induced by collisions actually dampens the oscillation (fig. 2b).



Figure 2 — Non linear Monte Carlo for gas kinetics.



**Fig. 2 Non-Linear Monte Carlo for gas kinetics.**

**a, Relaxation of speed distribution to equilibrium** (**a1**: tail distribution of particles speeds (fraction of particles faster than  $S$ ), **a2**: probability density function of velocity, **a3**: evolution of the fraction of particles faster than 6 speed units). In an homogenous gas, collisions between particles redistribute velocities so that the speed distribution tends to equilibrium (Maxwellian distribution). Starting from a distribution far from equilibrium (black curves), we compute the relaxation to the Maxwellian distribution (blue curves) by estimating the tail distribution at different times (e.g. red curves correspond to the system state at 1 unit time). The continuous lines correspond to the analytical solutions and each point corresponds to an independent NLMC estimation. The fraction of particles faster than 6 speed units (**a3**) illustrates how NLMC well accounts for the 1000-fold increase of the rarest high-speed particles, with no space nor time discretisation. Remarkably, rare events are estimated with the same relative uncertainty as frequent events ( $10^4$  samples for each estimate, confidence intervals of all estimates are contained within the point thickness).

**b, Dampening of breathing mode.** A dilute gas confined by an outside harmonic potential display a *breathing mode*. We estimate the density at probe position  $(1.75, 0, 0)$  and velocity  $(0.35, 0, 0)$  (in adimensional units, see SI5) at different times (adim. unit). Starting from a distribution complying with local equilibrium, the density displays a perpetual oscillation at twice the trap frequency (blue curve, analytical solution), independent of the collision rate (or, equivalently, of the elastic cross section). Starting from a distribution far from local equilibrium (the same initial distribution as in **a**), the density still pulsates in absence of collisions (null cross section, black curve: analytical solution, black points: probe estimates). Introducing collisions (raising cross section), the velocities redistribution induced by collisions dampens the oscillation (red points: probe estimates, no analytical solution). The inset illustrates that probes can be bunched to zoom in on a period of special interest (e.g. estimating the peak values at each cycle). Each point is estimated independently using  $10^7$  samples, confidence intervals of all estimates are contained within the point thickness in the main figure, confidence intervals are represented by the gray background area in the inset).

## Conclusions

From now on, the Monte Carlo Method is no more restricted to linear problems. The five examples exposed above have been worked out by teams gathering specialists of Monte Carlo method and specialists of the physical problem under consideration. By their full exposition, we offer the readers all the details to implement their own applications. As a guideline, the first step is to formulate the physical observable under its expectation form, including the non linearities and integrating all levels of complexity. The second step is to reformulate this expectation as a formulation compliant with standard Monte Carlo Method, according to the kind of non linearities. For polynomial non linearities, use i.i.d. series products. For other differentiable forms, use Taylor expansion around an upper-bound of the innermost random variable in order to regain a polynomial form. Using this MC-compliant formulation, every advanced MC technique can then be applied: parallel implementation, complex geometry, null-collisions, zero-variance, control variate, importance sampling, sensitivities analysis... As illustrated by the variety of our seminal examples, this guideline covers a large set of non linear academic and real world problems.

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## Methods

### Basics of Monte Carlo Methods.

Let us estimate  $E = 1 + 4$  by repeatedly tossing a (fair) coin. The tossing process is described by a random variable  $R \in \{0, 1\}$  that takes the value 0 for head (probability  $P_R(0) = \frac{1}{2}$ ) and 1 for tail (probability  $P_R(1) = \frac{1}{2}$ ).

Now, to estimate any process (e.g. a process output:  $E = 1 + 4$ ), we can affect arbitrary weights  $w(R)$  to values  $\{0, 1\}$  in order to write  $E$  as an expectation of the weighted process, following:

$$E = 1 + 4 = P_R(0)w(0) + P_R(1)w(1) = \mathcal{E}_R(w(R)) \quad (3)$$

with  $w(0) = \frac{1}{P_R(0)} = 2$  and  $w(1) = \frac{4}{P_R(1)} = 8$  and where  $\mathcal{E}_R$  denotes expectation with respect to  $R$ . Using the results  $r_1 \dots r_N$  of  $N$  successive tosses (independent realisations of  $R$ ), we can then estimate  $E = \mathcal{E}_R(w(R))$  from the weighted average of the tosses' results  $\frac{1}{N} \sum_{i=1}^N w(r_i)$  since  $E = 5$  is indeed the average of Monte Carlo weights that take the values 2 and 8 with equal probabilities.

Such an approach is at the basement of Monte Carlo techniques: define the weights according to the problem to be solved, sample the process repeatedly, and take the average. Depending on the physical description of the value to be estimated, this approach still holds for an infinite number of terms and can also be extended to integrals formulation using continuous random variables:

$$\mathcal{E}_Y(w(\mathbf{Y})) = \int_{\mathcal{D}_Y} d\mathbf{y} p_Y(\mathbf{y}) w(\mathbf{y}) \quad (4)$$

which can be estimated by  $\frac{1}{N} \sum_{i=1}^N w(\mathbf{y}_i)$ , where the  $y_i$  are  $N$  realisations of the random variable  $Y$  with probability density function  $p_Y$  and domain of definition  $\mathcal{D}_Y$ .

### Basics of Non-Linear Monte Carlo Methods.

In order to estimate

$$E = \mathcal{E}_Y(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})^2) = \int_{\mathcal{D}_Y} d\mathbf{y} p_Y(y) \left( \int_{\mathcal{D}_{\mathbf{X}|\mathbf{Y}}} dx p_{\mathbf{X}|\mathbf{Y}}(x|y) x \right)^2 \quad (5)$$

we introduce two independent variables  $\mathbf{X}_1$  and  $\mathbf{X}_2$  identically distributed as  $\mathbf{X}$  (still conditioned by the same  $\mathbf{Y}$ ):

$$\begin{aligned} E &= \mathcal{E}_Y(\mathcal{E}_{\mathbf{X}_1|\mathbf{Y}}(\mathbf{X}_1|\mathbf{Y}) \mathcal{E}_{\mathbf{X}_2|\mathbf{Y}}(\mathbf{X}_2|\mathbf{Y})) \\ &= \int_{\mathcal{D}_Y} d\mathbf{y} p_Y(y) \left( \int_{\mathcal{D}_{\mathbf{X}_1|\mathbf{Y}}} dx_1 p_{\mathbf{X}_1|\mathbf{Y}}(x_1|y) x_1 \right) \left( \int_{\mathcal{D}_{\mathbf{X}_2|\mathbf{Y}}} dx_2 p_{\mathbf{X}_2|\mathbf{Y}}(x_2|y) x_2 \right) \end{aligned} \quad (6)$$

Since  $\mathbf{X}_1$  and  $\mathbf{X}_2$  are independent, and conditionally independent given  $\mathbf{Y}$ :

$$\begin{aligned} E &= \int_{\mathcal{D}_Y} d\mathbf{y} p_Y(y) \left( \iint_{\mathcal{D}_{\mathbf{X}_1|\mathbf{Y}} \times \mathcal{D}_{\mathbf{X}_2|\mathbf{Y}}} dx_1 p_{\mathbf{X}_1|\mathbf{Y}}(x_1|y) dx_2 p_{\mathbf{X}_2|\mathbf{Y}}(x_2|y) x_1 x_2 \right) \\ &= \mathcal{E}_Y(\mathcal{E}_{(\mathbf{X}_1, \mathbf{X}_2)|\mathbf{Y}}(\mathbf{X}_1 \mathbf{X}_2|\mathbf{Y})) \end{aligned} \quad (7)$$

Hence

$$\begin{aligned} E &= \iiint_{\mathcal{D}_{\mathbf{Y}} \times \mathcal{D}_{\mathbf{X}|\mathbf{Y}}^2} dy p_{\mathbf{Y}}(y) dx_1 p_{\mathbf{X}|\mathbf{Y}}(x_1|y) dx_2 p_{\mathbf{X}|\mathbf{Y}}(x_2|y) x_1 x_2 \\ &= \mathcal{E}_{\mathbf{Y},(\mathbf{X}_1, \mathbf{X}_2)|\mathbf{Y}}(\mathbf{X}_1 \mathbf{X}_2 | \mathbf{Y}) \end{aligned} \quad (8)$$

The same demonstration can be drawn to establish that:

$$\mathcal{E}_{\mathbf{Y}}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})^n) = \mathcal{E}_{\mathbf{Y},(\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n)|\mathbf{Y}}(\mathbf{X}_1 \mathbf{X}_2 \dots \mathbf{X}_n | \mathbf{Y}) \quad (9)$$

Let us now assume that the weights associated to the random variable  $\mathbf{Y}$  are described by a nonlinear function  $\mathbf{f}(\mathbf{Z}_{\mathbf{Y}})$  of the conditional expectation  $\mathbf{Z}_{\mathbf{Y}} = \mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$ . The problem then becomes to compute:

$$E = \mathcal{E}_{\mathbf{Y}}(\mathbf{f}(\mathbf{Z}_{\mathbf{Y}})) = \mathcal{E}_{\mathbf{Y}}(\mathbf{f}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y}))) \quad (10)$$

Such nonlinearity can be treated with no approximation using a projection on infinite basis. In all the examples presented in this article, we have used Taylor polynomials basis, which means that  $\mathbf{f}(x)$  is expanded around  $x_0$

$$\mathbf{f}(x) = \sum_{n=0}^{+\infty} \frac{\partial^n \mathbf{f}(x_0)}{n!} (x - x_0)^n \quad (11)$$

We note that both  $x_0$  and  $\mathbf{f}$  can be conditioned by  $\mathbf{Y}$ . Now, following the same line as explained above for the Basics of Monte Carlo Methods, we regard the sum in the expansion of  $\mathbf{f}$  as an expectation, writing:

$$\mathbf{f}(x) = \mathcal{E}_N \left( \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} (x - x_0)^N \right) \quad (12)$$

where the random variable  $N$  (of probability law  $P_N$ ) is the degree of one monomial in the Taylor polynomial. This step only requires to define an infinite set of probabilities (instead of two in Eq. 3), with  $\sum_{n=0}^{+\infty} P_N(n) = 1$ . For the choice of  $x_0$ , we have used an upper bound of  $\mathbf{X}|\mathbf{Y}$ , so that  $P_N$  is implicitly derived using Bernoulli random variables (see details in SI 1-3).

Equation 10 can then be rewritten as:

$$E = \mathcal{E}_{\mathbf{Y}}(\mathbf{f}(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y}))) = \mathcal{E}_{Y,N} \left( \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} (\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y}) - x_0)^N \right) \quad (13)$$

Defining independent and identically distributed random variables  $\mathbf{X}_q$ , with same distribution as  $\mathbf{X}$ , the innermost term rewrites

$$E = \mathcal{E}_{Y,N} \left( \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} \prod_{q=1}^N (\mathcal{E}_{\mathbf{X}_q|\mathbf{Y}}(\mathbf{X}_q|\mathbf{Y}) - x_0) \right) \quad (14)$$

or, equivalently:

$$E = \mathcal{E}_{Y,N} \left( \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} \prod_{q=1}^N \mathcal{E}_{\mathbf{X}_q|\mathbf{Y}}(\mathbf{X}_q|\mathbf{Y} - x_0) \right) \quad (15)$$

Since the variables  $\mathbf{X}_q|\mathbf{Y}$  are independent in the innermost term, we have:

$$\prod_{q=1}^N \mathcal{E}_{\mathbf{X}_q|\mathbf{Y}}(\mathbf{X}_q|\mathbf{Y} - x_0) = \mathcal{E}_{(\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N)|\mathbf{Y}} \left( \prod_{q=1}^N (\mathbf{X}_q|\mathbf{Y} - x_0) \right) \quad (16)$$

so that:

$$E = \mathcal{E}_{\mathbf{Y}, \mathbf{N}} \left( \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} \mathcal{E}_{(\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N) | \mathbf{Y}} \left( \prod_{q=1}^N (\mathbf{X}_q | \mathbf{Y} - x_0) \right) \right) \quad (17)$$

and we finally have:

$$E = \mathcal{E}_{\mathbf{Y}, \mathbf{N}, (\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N) | \mathbf{Y}} \left( \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} \prod_{q=1}^N (\mathbf{X}_q | \mathbf{Y} - x_0) \right) \quad (18)$$

which can be read as:

$$E = \mathcal{E}_{\mathbf{Y}, \mathbf{N}, (\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N) | \mathbf{Y}} (w(\mathbf{Y}, \mathbf{N}, \mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N)) \quad (19)$$

with

$$w(\mathbf{Y}, \mathbf{N}, \mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_N) = \frac{\partial^N \mathbf{f}(x_0)}{P_N(N)N!} \prod_{q=1}^N (\mathbf{X}_q | \mathbf{Y} - x_0) \quad (20)$$

The translation into a Monte Carlo algorithm then follows:

- sample a realisation  $\mathbf{y}$  of  $\mathbf{Y}$  (and set  $x_0$  and  $\mathbf{f}$  accordingly if they depend on  $\mathbf{y}$ )
- sample a realisation  $n$  of  $N$
- sample  $n$  independent realisations  $\mathbf{x}_{q=1, \dots, n}$  of  $\mathbf{X}$  conditioned by  $\mathbf{y}$
- keep

$$\hat{w} = w(\mathbf{y}, n, \mathbf{x}_1, \dots, \mathbf{x}_n) = \frac{\partial^n \mathbf{f}(x_0)}{P_N(n)n!} \prod_{q=1}^n (\mathbf{x}_q - x_0)$$

and estimate  $E$  as the average of many realisations  $\hat{w}$ .



## Supplemental Information

# 1 SI1 — Solar thermochemical reduction of zinc oxide averaged over the year

## 1.1 The problem

Solar-driven high temperature thermochemical cycles processes, commonly based on metal oxides reduction [1, 2], are an alternative to fossil fuel-based method for  $H_2$  generation. Here we focus on thermal reduction zinc oxide, as the first part of a two step water splitting cycle. Photons emitted from the sun are reflected on heliostats and concentrated at the entrance of the chemical reactor in which ZnO dissociation is carried out. Solar power received by the reactor at a given instant determines the chemical conversion rate of the reaction  $ZnO \rightarrow Zn + \frac{1}{2}O$ . The present Monte Carlo estimation of the solar-plant's annual conversion rate  $C$  associates the thermochemical knowledge of the non-linear kinetics of zinc oxide dissociation [3, 4] to the description of radiative transfer on a multiple-reflection solar receiver [5]. The nonlinearity lies in the instantaneous coupling between photon transport and zinc-oxide reduction.

The random variable  $\mathbf{X}$  is the contribution of an optical path from the sun to the entrance of the chemical reactor. Its expectation  $\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})$  is the instantaneous thermal-power fraction collected at a given moment  $\mathbf{Y}$  of the year[6]. Solar power is used to reduce the zinc oxide with a nonlinear conversion rate  $f(\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y}))$  (see[3, 4]).

## 1.2 Non-Linear Monte Carlo formulation

The annual conversion rate  $C$  is reformulated as its Taylor expansion around  $x_0$ . First, independent and identically distributed (i.i.d) optical-path random variables are introduced, each defining a random contribution  $\mathbf{X}_i|\mathbf{Y}$  i.i.d as  $\mathbf{X}|\mathbf{Y}$ . Then, we expand the non-linear function  $f$  around  $x_0$  chosen as an upper bound of  $\mathbf{X}|\mathbf{Y}$ . Finally, the infinite power series is statistically formulated thanks to the discrete random variable  $N$  the order of Taylor expansion:

$$N_{\mathbf{Y}} = B_{\mathbf{X}_1|\mathbf{Y}} + \sum_{i=2}^{+\infty} i B_{\mathbf{X}_i|\mathbf{Y}} \prod_{q=1}^{i-1} (1 - B_{\mathbf{X}_q|\mathbf{Y}}) \quad (21)$$

where the  $B_{\mathbf{X}_i|\mathbf{Y}}$  are Bernoulli random variables:

$$B_{\mathbf{X}_i|\mathbf{Y}} = \begin{cases} 1 & \text{with probability } P_{\mathbf{X}_i|\mathbf{Y}} = \frac{\mathbf{X}_i|\mathbf{Y}}{x_0} \\ 0 & \text{with probability } 1 - P_{\mathbf{X}_i|\mathbf{Y}} \end{cases} \quad (22)$$

In the end, the reformulation is:

$$C = \mathcal{E}_{\mathbf{Y},(\mathbf{X}_1,\mathbf{X}_2,\dots,\mathbf{X}_{N_{\mathbf{Y}}})|\mathbf{Y}}(w(N_{\mathbf{Y}})) \quad (23)$$

where the Monte Carlo weight function  $w$  is

$$w(u) = \begin{cases} \exp(C_4 x_0^{-\alpha}) \sum_{m=1}^3 C_m x_0^{(m\alpha-1)} & \text{if } u = 1 \\ w(1) + \exp(C_4 x_0^{-\alpha}) \sum_{k=0}^u (-1)^k \sum_{p=0}^k \frac{1}{p!(k-p)!} \sum_{m=1}^3 C_m x_0^{(m\alpha-1+p)} \\ \quad \times \prod_{j=1}^{k-p} (m\alpha - j) \sum_{r=1}^p (-1)^r C_4 x_0^{-(\alpha+r)} \prod_{l=0}^{r-1} (\alpha + l) & \text{if } u > 1 \end{cases} \quad (24)$$

with  $u \in \mathbb{N}$  and  $C_1, C_2, C_3, C_4, \alpha$  and  $x_0$  known constants.

### 1.3 Algorithm

**Step 1** Uniform sampling of a moment  $\mathbf{y}$  of the year and index initialisation  $n = 1$ .

**Step 2** Sampling of the  $n$ -th optical path contributing to thermal-power collection at instant  $\mathbf{y}$  and computation of its contribution  $x_n$  (as presented in [5, 6]).

**Step 3** Computation of the probability  $P_n$  in Eq. 22 and sampling of a realisation  $b_n$  of the Bernoulli random variable  $B_{\mathbf{X}_n|\mathbf{Y}}$ .

**Step 4** If  $b_n = 0$ , the algorithm loops to step 2 after incrementation of  $n$ . Else, the procedure is terminated and the weight  $\hat{w}(n)$  computed according to Eq. 24.

### 1.4 Simulated configuration

Results shown in Extended Data Figure 1 are obtained for a 1 MW solar plant with a 80 m high central receiver tower and 160 heliostats arranged in a radial staggered layout (nueen method). Values of the conversion-rate parameters  $C_1, C_2, C_3, C_4$  and  $\alpha$  are given in [3, 4]. Meteonorm DNI database.

## 2 SI2 — Wave scattering

This example is fully detailed in Charon *et al.* 2016[7].

### 2.1 The problem

Here we address the solution of Schiff's approximation[8], also known as the *anomalous diffraction approximation*[9], for an incident plane wave with propagation direction  $\mathbf{e}_i$  and wave number  $k$  scattered by the scattering potential  $U$  that takes value  $U^{in}$  inside a large domain compared to  $1/k$  and 0 outside (this domain corresponds to the shape of the scatterer), with  $|U^{in}| \ll k^2$ . Let  $\Gamma_{\mathbf{Y}}$  be the projected surface of the scatterer seen from  $\mathbf{e}_i$ , given the scatterer orientation  $\mathbf{Y}$ . Straight rays along  $\mathbf{e}_i$  cross  $\Gamma_{\mathbf{Y}}$  at locations  $\mathbf{P}|\mathbf{Y}$ ; the domain of definition of  $\mathbf{P}|\mathbf{Y}$  is then  $\mathcal{D}_{\mathbf{P}|\mathbf{Y}} \equiv \mathbf{P}|\mathbf{Y}$ . These rays are attenuated and phase shifted over the crossing length  $l(\mathbf{P}|\mathbf{Y})$ . The complex random variable  $X|\mathbf{Y} = X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s)$  is the contribution of one of these rays to the scattered field in a direction  $\mathbf{e}_s$  deviating from  $\mathbf{e}_i$  by a small scattering angles  $\theta$ [7]:

$$X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s) = A_{\mathbf{Y}} \frac{k}{2\pi} \exp(ik\theta \mathbf{b} \cdot \mathbf{P}|\mathbf{Y}) \left\{ 1 - \exp\left(\frac{i}{2k} U^{in} l(\mathbf{P}|\mathbf{Y})\right) \right\} \quad (25)$$

where  $\mathbf{b}$  is the unit vector along the projection of  $\mathbf{e}_s$  on the plane containing  $\mathcal{D}_{\mathbf{P}|\mathbf{Y}}$  and  $A_{\mathbf{Y}}$  is the area of  $\mathcal{D}_{\mathbf{P}|\mathbf{Y}}$ . Given the orientation  $\mathbf{Y}$  of the scatterer, the conditional expectation  $S_{\mathbf{Y}}(\mathbf{e}_s) = \mathcal{E}_{X|\mathbf{Y}}(X|\mathbf{Y}) = \mathcal{E}_{\mathbf{P}|\mathbf{Y}}(X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s))$  is the complex scattering amplitude in direction  $\mathbf{e}_s$ [8, 9, 10, 7]. For the scatterer orientation  $\mathbf{Y}$ , the far-field scattering diagram is given by the *differential scattering cross-section*[9, 12]:  $\hat{W}_{\mathbf{Y}}(\mathbf{e}_s) = |S_{\mathbf{Y}}(\mathbf{e}_s)|^2 = |\mathcal{E}_{\mathbf{P}|\mathbf{Y}}(X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s))|^2$ .

We address the Monte Carlo computation of  $W(\mathbf{e}_s) = \mathcal{E}_{\mathbf{Y}}(\hat{W}_{\mathbf{Y}}(\mathbf{e}_s))$  which is  $\hat{W}_{\mathbf{Y}}(\mathbf{e}_s)$  averaged over the statistics of the scatterer orientation  $\mathbf{Y}$  [9, 12, 10]. The full expression is then

$$\begin{aligned} W(\mathbf{e}_s) &= \mathcal{E}_{\mathbf{Y}}(|S_{\mathbf{Y}}(\mathbf{e}_s)|^2) \\ &= \mathcal{E}_{\mathbf{Y}}(|\mathcal{E}_{\mathbf{P}|\mathbf{Y}}(X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s))|^2) \\ &= \mathcal{E}_{\mathbf{Y}}\left((\Re \mathcal{E}_{\mathbf{P}|\mathbf{Y}}(X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s)))^2 + (\Im \mathcal{E}_{\mathbf{P}|\mathbf{Y}}(X_{\mathbf{P}|\mathbf{Y}}(\mathbf{e}_s)))^2\right) \end{aligned} \quad (26)$$

where the configuration spaces are the orientation vectors of the scattering potential with respect to  $\mathbf{e}_i$  ( $\mathcal{D}_{\mathbf{Y}}$ ) and the projected surface of the scattering potential seen from the incident direction  $\mathbf{e}_i$  ( $\mathcal{D}_{\mathbf{P}|\mathbf{Y}}$ ).

### 2.2 Non-Linear Monte Carlo formulation

$W(\mathbf{e}_s)$  is reformulated based on the definition of two independent and identically distributed location random-variables  $\mathbf{P}_1|\mathbf{Y}$  and  $\mathbf{P}_2|\mathbf{Y}$ :

$$\begin{aligned} W(\mathbf{e}_s) &= \mathcal{E}_{\mathbf{Y}}(\Re \mathcal{E}_{\mathbf{P}_1|\mathbf{Y}}(X_{\mathbf{P}_1|\mathbf{Y}}(\mathbf{e}_s)) \Re \mathcal{E}_{\mathbf{P}_2|\mathbf{Y}}(X_{\mathbf{P}_2|\mathbf{Y}}(\mathbf{e}_s)) + \Im \mathcal{E}_{\mathbf{P}_1|\mathbf{Y}}(X_{\mathbf{P}_1|\mathbf{Y}}(\mathbf{e}_s)) \Im \mathcal{E}_{\mathbf{P}_2|\mathbf{Y}}(X_{\mathbf{P}_2|\mathbf{Y}}(\mathbf{e}_s))) \\ &= \mathcal{E}_{\mathbf{Y},(\mathbf{P}_1,\mathbf{P}_2)|\mathbf{Y}}(w(\mathbf{P}_1, \mathbf{P}_2, \mathbf{e}_s)) \end{aligned} \quad (27)$$

where the Monte Carlo weight function  $w$  is:

$$w(\mathbf{P}_1, \mathbf{P}_2, \mathbf{e}_s) = \Re X_{\mathbf{P}_1|\mathbf{Y}}(\mathbf{e}_s) \Re X_{\mathbf{P}_2|\mathbf{Y}}(\mathbf{e}_s) + \Im X_{\mathbf{P}_1|\mathbf{Y}}(\mathbf{e}_s) \Im X_{\mathbf{P}_2|\mathbf{Y}}(\mathbf{e}_s) \quad (28)$$

with  $X_{\mathbf{P}_q|\mathbf{Y}}(\mathbf{e}_s)$  defined in Eq. 25.

## 2.3 Algorithm

The sampling procedures of the Monte-Carlo algorithm are then:

**Step 1** Isotropic sampling of an orientation  $\mathbf{y}$  of the scattering potential.

**Step 2** Uniform sampling of the first location  $\mathbf{p}_1$  on the projected surface defined by  $\mathbf{y}$  and computation of the corresponding crossing length  $l(\mathbf{p}_1)$ : the realisation  $x_1$  of  $X_{\mathbf{p}|\mathbf{Y}}(\mathbf{e}_s)$  is computed according to Eq. 25.

**Step 3** Uniform sampling of the second location  $\mathbf{p}_2$  on the projected surface defined by  $\mathbf{y}$  and computation of the corresponding crossing length  $l(\mathbf{p}_2)$ : the realisation  $x_2$  of  $X_{\mathbf{p}|\mathbf{Y}}(\mathbf{e}_s)$  is computed according to Eq. 25.

**Step 4** Computation of the weight  $\hat{w} = w(\mathbf{p}_1, \mathbf{p}_2, \mathbf{e}_s)$  according to Eq. 28:  $\hat{w} = \Re x_1 \Re x_2 + \Im x_1 \Im x_2$

Codes for the implementation of this algorithm in the case of spheroidal and cylindrical scattering potentials, as well as validation against reference solution, are provided in Charon et al. 2016 [7] (see also <http://edstar.lmd.jussieu.fr/codes>).

## 2.4 Simulated configuration

Results shown in Extended Data Figure 2 are obtained for  $k = 2\pi/\lambda$ , with wavelength  $\lambda = 500nm$  and the scattering potential  $U^{in} = -2k(m-1)$ , with  $m = 1.2 - i 4 \cdot 10^{-3}$  (scatterer relative refractive index[10]). The shape of the scatterer is helical with length  $L = 50\mu m$ , pitch  $P = 15\mu m$ , helix diameter  $D = 20\mu m$ , cylinder diameter  $d = 3.5\mu m$ .

### 3 SI3 — Phytoplankton growth in light-limited environments

#### 3.1 The problem

Here we address the production of reference solutions for a photobioreactor model. This model is based on a radiative transfer approach presented in[11]. The rate of photon absorption  $\mathbf{A}_Y$  by micro-algae at location  $Y$  within the culture volume is solution of the Radiative Transfer Equation. The present study is based on the standard linear transport MC algorithm presented in[15] for the estimation of  $\mathbf{A}_Y$  at any location within the photobioreactor. In this algorithm, realisations of the optical path random variable  $\Gamma(C)|Y$  are sampled backward from the absorption location  $Y$  to the light emitting surface (the surface of the 979 light-diffusing optical fibres, see EDF3). Since micro-algae are scattering visible light,  $\Gamma(C)|Y$  depends on their concentration  $C$  within the suspension: the scattering and absorption coefficient of the suspension are proportional to  $C$  (independent scattering). The random variable  $\mathbf{X}(C)|Y = \mathbf{X}_{\Gamma(C)|Y}(C)$  is the contribution of one of these optical paths to the photon absorption rate at  $Y$  (see[15] for the detailed expression of  $\mathbf{X}_{\Gamma(C)|Y}(C)$ ). This contribution  $\mathbf{X}_{\Gamma(C)|Y}(C)$  depends on the biomass concentration  $C$  because the absorption coefficient of the suspension determines the fraction of incident light flux that is transmitted along the optical path  $\Gamma(C)|Y$  (according to Beer law). In the end, the conditional expectation  $\mathbf{A}_Y(C) = \mathcal{E}_{\mathbf{X}|Y}(\mathbf{X}|Y) = \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))$  is the rate of photon absorption at location  $Y$ , for the biomass concentration  $C$ .

Photons absorbed by a micro-algae are converted within the photosynthetic units and the Z-scheme[17, 18], leading to the local biomass growth-rate[14, 11]

$$\begin{aligned} r_Y(C) &= f(\mathbf{A}_Y(C), C) \\ &= C \left( \alpha \frac{K \mathbf{A}_Y(C)}{K + \mathbf{A}_Y(C)} - \beta \frac{K_r}{K_r + \mathbf{A}_Y(C)} \right) \\ &= C \left( \alpha \frac{K \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))}{K + \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))} - \beta \frac{K_r}{K_r + \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))} \right) \end{aligned} \quad (29)$$

where  $\alpha$ ,  $\beta$ ,  $K$  and  $K_r$  are constant parameters that depend on the studied microorganism (see[14] for values in the case of *Chlamydomonas Reinhardtii*).

Due to light absorption and scattering by micro-algae, the field  $\mathbf{A}_Y$  is heterogeneous within the volume (less light farther from the fibres and for higher concentrations) and so is the local photosynthetic rate  $r_Y(C)$ . We address the Monte Carlo computation of  $R(C) = \mathcal{E}_Y(r_Y(C))$ , the local photosynthetic rate  $r_Y(C)$  averaged over locations  $Y$  of the culture volume. The full expression is then

$$R(C) = \mathcal{E}_Y \left( C \left( \alpha \frac{K \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))}{K + \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))} - \beta \frac{K_r}{K_r + \mathcal{E}_{\Gamma(C)|Y}(\mathbf{X}_{\Gamma(C)|Y}(C))} \right) \right) \quad (30)$$

#### 3.2 Non-Linear Monte Carlo formulation

$R(C)$  is reformulated based on the Taylor expansion of  $f$  around  $x_0$ . First, independent and equally distributed optical-path random variables  $\Gamma_i(C)|Y$  are introduced, defining independent and equally distributed contributions  $\mathbf{X}_i(C)|Y = \mathbf{X}_{\Gamma_i(C)|Y}(C)$ . Then, we expand the non-linear function in Eq. 29 around  $x_0$  chosen as an upper bound of  $\mathbf{X}_i(C)|Y$ . Finally, the infinite power series is statistically formulated thanks to the discrete random variable  $N$  the order of Taylor expansion:

$$N_Y(C) = B_{\mathbf{X}_1(C)|Y} + \sum_{i=2}^{+\infty} i B_{\mathbf{X}_i(C)|Y} \prod_{q=1}^{i-1} (1 - B_{\mathbf{X}_q(C)|Y}) \quad (31)$$

where the  $B_{\mathbf{X}_i(C)|\mathbf{Y}}$  are Bernoulli random variables:

$$B_{\mathbf{X}_i(C)|\mathbf{Y}} = \begin{cases} 1 & \text{with probability } P_{\mathbf{X}_i(C)|\mathbf{Y}} = \frac{\mathbf{X}_i(C)|\mathbf{Y}}{x_0} \\ 0 & \text{with probability } 1 - P_{\mathbf{X}_i(C)|\mathbf{Y}} \end{cases} \quad (32)$$

In the end, the reformulation is:

$$R(C) = \mathcal{E}_{\mathbf{Y}, (\mathbf{r}_1, \mathbf{r}_2, \dots, \mathbf{r}_{N_{\mathbf{Y}}})|\mathbf{Y}} (w(N_{\mathbf{Y}})) \quad (33)$$

where the Monte Carlo weight function  $w$  is

$$w(u) = \begin{cases} C \left[ \alpha \frac{K x_0}{K + x_0} - \beta \frac{K_r}{K_r + x_0} \right] & \text{if } u = 1 \\ w(1) - \sum_{i=1}^{u-1} C \left[ \alpha \frac{K^2}{K + x_0} \left( \frac{x_0}{K + x_0} \right)^i + \beta \frac{K_r}{K_r + x_0} \left( \frac{x_0}{K + x_0} \right)^i \right] & \text{if } u > 1 \end{cases} \quad (34)$$

with  $u \in \mathbb{N}$  and  $\alpha, K, \beta, K_r, x_0$  and  $C$  known constants.

### 3.3 Algorithm

**Step 1** Uniform sampling of a location  $\mathbf{y}$  within the culture volume and index initialisation  $n = 1$ .

**Step 2** Sampling of the  $n$ -th optical path contributing to absorption at  $\mathbf{y}$ : the realisation  $x_n$  of  $\mathbf{X}_{\mathbf{r}_n(C)|\mathbf{Y}}(C)$  is computed according to [15].

**Step 3** Computation of the probability in Eq. 32,  $P = x_n/x_0$ , and sampling of a realisation  $b_n$  of the Bernoulli random variable  $B_{\mathbf{X}_n(C)|\mathbf{Y}}$ .

**Step 4** If  $b_n = 0$ , the algorithm loops to step 2 after incrementation of  $n$ . Else, the procedure is terminated and the weight  $\mathbf{w}(n)$  computed according to Eq. 34.

### 3.4 Simulated configuration

Results shown in Extended Data Figure 3 are obtained for kinetics parameters  $\alpha = 1.785 \cdot 10^{-9} \text{ kg}/\mu\text{mol}_{h\nu}$ ,  $\beta = 4.057 \cdot 10^{-6} \text{ s}^{-1}$ ,  $K = 32000 \mu\text{mol}_{h\nu}/\text{kg}/\text{s}$  and  $K_r = 7500 \mu\text{mol}_{h\nu}/\text{kg}/\text{s}$ . The photobioreactor is a 25L DiCoFluV[13]: reactor diameter 16.5 cm, reactor height 1 m, optical fibres diameter 1.2 mm, fibres height 1 m, 979 fibres, hexagonal lattice fibre arrangement with distance 4.8 mm between two fibres axis (see [15, 13]), homogeneous surface flux density  $25 \mu\text{mol}_{h\nu}/\text{m}^2/\text{s}$  emitted at fibres surface. The radiative properties (absorption and scattering by micro-algae) are those presented in [10] for *Chlamydomonas Reinhardtii*.

## 4 SI4 — Atmospheric radiative transfer: top-of-atmosphere specific intensity (from earth toward the outer space)

This example is fully detailed in Galtier *et al.* 2015[19].

### 4.1 The problem

Photons are emitted either by a surface (here the ground) or by the volume (the atmosphere). The radiation  $I(\Delta\nu)$  perceived at the observation location TOA is the fraction of all these photons that have a frequency inside the observation-band  $\Delta\nu$  and reach the sensor, i.e. which are not absorbed by the atmosphere in between. When no scattering occurs (clear sky), photon-paths are straight lines and the fraction of the photons of frequency  $\nu_y$  that travel (let say along the strict vertical) from altitude  $H_y$  to TOA is

$$\exp(-\tau_y)$$

where

$$\tau_y = \int_{H_y}^{TOA} dx \sum_{j_x=1}^{N_t} h_{\nu_y, j_x}(x)$$

This exponential extinction is that of Beer law where  $\tau_y$  is the monochromatic optical thickness, i.e. the sum of all transition cross-sections  $h_{\nu_y, j_x}(x)$  at all intermediate altitudes  $x$ . In this example, the random variable  $\mathbf{Y}$  is a vector that gathers the altitude  $H_{\mathbf{Y}}$  of emission, the photon-path  $\Gamma_{\mathbf{Y}}$  from  $H_{\mathbf{Y}}$  to TOA and frequency  $\nu_{\mathbf{Y}}$  of emission. Similarly,  $\mathbf{X}$  gathers all the description of an absorption event, altitude  $H_{\mathbf{X}}$  and index of absorption-transition  $j_{\mathbf{X}}$ .

In standard Monte Carlo approaches,  $\tau_y$  is either precomputed and tabulated, or is easily computable from tabulated values of the absorption coefficient ( $k = \sum_{j_x=1}^{N_t} h_{\nu_y, j_x}$ ). Then the Monte Carlo algorithm deals only with the sampling of  $\mathbf{Y}$  and the weight function has the form

$$w(y) = \mathcal{S}_y \exp(-\tau_y)$$

$\mathcal{S}_{\mathbf{Y}}$  is the source associated with the emission  $\mathbf{Y}$  (Planck's function at the local temperature times an emissive power that depends on concentration, pressure and temperature). The sum over the  $N_t$  transitions  $j_x$  is thus not handled by standard Monte Carlo algorithms, despite the fact that this sum has all the features inviting to make use of statistical approaches:  $N_t$  is huge, typically of the order of  $10^6$ , and the deterministic pre-calculations can be computationally very demanding (and are to be re-performed when testing each new spectroscopic assumption). No attempt has been made so far to address this sum statistically, together with the photon-path statistics because these two statistics are combined via the non-linearity of the exponential extinction.

### 4.2 Non-Linear Monte Carlo formulation

So the present question is to design a Monte Carlo algorithm performing the summation  $\sum_{j=1}^{N_t} h_{\nu, j}$  together with the integrals over frequency, emission location and absorption location, despite of the nonlinearity of Beer extinction. This starts by looking at  $\tau_y$  in statistical terms : the integral over altitudes  $H_{\mathbf{X}}$  is transformed into the sampling of absorption locations, and the sum  $\sum_{j_x=1}^{N_t} h_{\nu_y, j_x}$  into the sampling of absorption-transitions. This double sampling is formally summarised into the random variable  $\mathbf{X}$ . To get a sample  $x$  of  $\mathbf{X}$ , one first samples  $H_{\mathbf{X}}$  and  $j_{\mathbf{X}}$ . Then  $x$  is the value



that the optical thickness  $\tau_y$  would have if the atmosphere was homogeneous at the thermodynamic conditions of altitude  $H_x$  and if all transitions were identical to  $j_x$ . This random variable depends on  $\mathbf{Y}$  and its expectation, knowing  $\mathbf{Y} = y$ , is  $\mathcal{E}(\mathbf{X}|\mathbf{Y} = y) = \tau_y$ . We therefore get  $I(\Delta\nu) = \mathcal{E}_{\mathbf{Y}}(\mathcal{S}_{\mathbf{Y}} \exp(-\mathcal{E}_{\mathbf{X}|\mathbf{Y}}(\mathbf{X}|\mathbf{Y})))$

### 4.3 Algorithm

A strict application of the solution described in *Methods* implies to first sample  $\mathbf{Y}$  (as in a standard Monte Carlo algorithm), then sample the degree  $n$  of one monomial in the Taylor expansion of the negative-exponential function  $f$ , and draw  $n$  independent samples  $x_1, x_2 \dots x_n$  of  $\mathbf{X}$ , i.e  $n$  absorption locations  $H_{x_q}$  (between emission location and TOA) and  $n$  transition indexes  $j_{x_q}$ :

**Step 1** sample a frequency,

**Step 2** sample an emission-altitude,

**Step 3** sample an order  $n$  of the development of the exponential,

**Step 4** sample  $n$  paired absorption-altitudes, transitions.

### 4.4 Null-collision reformulation

Instead of implementing this solution in this straightforward manner, we chose to use null-collisions[20]. The same quantities are sampled, but the order is different, leading to quite intuitive physical pictures. The first step is still to sample a frequency, but then we implement a backward tracking algorithm, starting from the observation altitude. We sample a first collision-altitude as if photons were emitted at the observation location in the downward direction in an homogeneous atmosphere. Then we sample a transition and a statistical test is made to determine whether the algorithm stops at this altitude or continues (a Bernoulli test to determine whether the collision is a true one or a null-collision). If it continues, this first collision-altitude and first transition have the status of one of the  $n$  absorption-altitudes and transitions in the above presented algorithm. Then a next collision-location is sampled, together with a next transition, etc. When the algorithm stops after  $n$  collisions, the final altitude is interpreted as the emission-altitude. So  $n$  and the emission-altitude are not sampled first: they are sampled by the successive Bernoulli tests as in a standard backward-tracking multiple-scattering algorithm.

### 4.5 The final algorithm

**Step 1** Initialisation of the current altitude at  $H_x \leftarrow TOA$ .

**Step 2** Uniform sampling of a frequency  $\nu_y$  in the considered infrared band (or the whole infrared).

**Step 3** Exponential sampling of a travelled distance  $d$  before absorption for a photon of frequency  $\nu_y$  travelling from  $H_x$  in the backward direction within a virtual homogeneous atmosphere of absorption coefficient  $\hat{k}_{\nu_y}$ .

**Step 4** If the travelled distance leads the photon to hit the surface,  $H_x \leftarrow 0$  and the algorithm jumps to *Step 7*. Otherwise  $H_x \leftarrow H_x - d$ .

**Step 5** Sampling of a state-transition  $j_x$  according to  $\mathcal{P}(1), \mathcal{P}(2) \dots \mathcal{P}(N_t)$  (see[19]).

**Step 6** Bernoulli trial of probability  $P_a = \frac{h_{\nu_y, j_x}}{k_{\nu_y} \mathcal{P}(j_x)}$  to decide whether the algorithm jumps to *Step 7* or *Step 3*.

**Step 7**  $H_y \leftarrow H_x$  and the algorithm stops with  $\hat{w} = \Delta\nu B_y$  where  $B_y$  is the Planck function at frequency  $\nu_y$  for the atmospheric temperature at  $H_y$ .

## 4.6 Simulated configuration

The simulation results in Extended Data Fig. 4 have been obtained using the HITRAN spectroscopic database. All other parameters are given in Galtier *et al.* 2015[19].

## 4.7 Computational performance

As far as computational costs are concerned, our computations did not require to first scan HITRAN and pre-compute absorption coefficients at all altitudes and all frequencies before running the Monte Carlo code for transfer since our Monte Carlo code handles both simultaneously despite the nonlinearity of the exponential extinction. The computational benefit is very significant when studying the effects of new spectroscopic data or new line-shape assumptions. Otherwise, for earth applications and fixed spectroscopic assumptions, there is no problem associated to the pre-computation of absorption coefficients and to their tabulation as function of molecular composition and thermodynamic state. But for combustion or astrophysics applications, thinking in particular of exoplanets, the diversity of compositions is extremely wide, temperatures can be high, imply the use of much larger spectroscopic databases, including hot lines, and pre-computation/tabulation of absorption coefficients is a today's challenge in itself. Our nonlinear Monte Carlo suppresses this need[19].

## 5 SI5 — Gas kinetics

### 5.1 The problem

We consider a gas of interacting particles, with collisions following Maxwell model, and a cross section  $\sigma(\vec{c}, \vec{c}^*) = \frac{\kappa}{4\pi\|\vec{c}-\vec{c}^*\|}$ . Particles are confined within an harmonic static trap of pulsation  $\omega$  so that acceleration at position  $\vec{r}$  is spring-like:  $\vec{a}(\vec{r}) = -\omega^2\vec{r}$ .

We follow the distribution function  $f(\vec{r}, \vec{c}, t)$  at location  $\vec{r}$ , velocity  $\vec{c}$  at time  $t$ . We consider it known at some time  $t_I$ , either constrained at local equilibrium, according to:

$$f_{\text{LEQ}}/n_I = p_{\mathcal{N}(\vec{u}_I, c_{q,I}^2)} \quad (35)$$

or, out of equilibrium, according to:

$$f_{\text{BKW}}/n_I = \frac{1}{3} \frac{(\vec{c} - \vec{u}_I)^2}{\frac{c_{q,I}^2}{5/3}} p_{\mathcal{N}(\vec{u}_I, \frac{c_{q,I}^2}{5/3})} \quad (36)$$

where  $n_I$  denotes density,  $\vec{u}_I$  denotes the mean velocity and  $c_{q,I}$  denotes the mean square speed at  $\vec{r}$  at time  $t_I$ ;  $p_{\mathcal{N}(\mu, \sigma^2)}$  denotes probability density function of a Gaussian random variable, with expectation  $\mu$  and variance  $\sigma^2$ .

**For the case in fig. 3a,** we set  $\omega = 0$ , and  $n_I$ ,  $\vec{u}_I$  and  $c_{q,I}$  are set homogeneous, corresponding to the case studied par Krook and Wu[21, 22]. Starting from  $f_{\text{BKW}}$ , the gas relaxes to equilibrium  $f_{\text{LEQ}}$ .

Adimensional time is  $\kappa n_I t$  and adimensional distribution function is  $\frac{1}{n_I} (\sqrt{2\pi} c_{q,I})^3 f$ .

**For the case in fig. 3b,** we set  $\omega = 2\pi$ , and  $n_I$ ,  $\vec{u}_I$  and  $c_{q,I}$  are set heterogeneous to correspond to one state of the undamped oscillation (*breathing mode*[23]):

$$\begin{aligned} n_I &= p_{\mathcal{N}(\vec{0}, \frac{c_{q,EQ}^2}{\omega^2})} \\ \vec{u}_I &= \epsilon \omega \vec{r} \\ c_{q,I} &= \sqrt{1 - \epsilon^2} c_{q,EQ} \end{aligned} \quad (37)$$

with  $\epsilon = \frac{\Delta c_q^2}{c_{q,EQ}^2}$  where  $\Delta c_q^2$  is the maximal deviation of  $c_q^2$  from its equilibrium value  $c_{q,EQ}^2$ .

With these initial values, and starting from local equilibrium  $f_{\text{LEQ}}$ , the local equilibrium remains unbroken and the gas displays the undamped oscillation for any value of the cross section. Conversely, starting from out of equilibrium  $f_{\text{BKW}}$ , any positive value of cross section will lead to dampened oscillations.

Adimensional time is  $\frac{\omega}{2\pi} t$ , adimensional distribution function is  $\left(2\pi \frac{c_{q,EQ}^2}{\omega}\right)^3 f$  and adimensional parameter for cross-section is  $\frac{\omega^2}{c_{q,EQ}^3} \kappa$ .

Simulation results are given for  $\epsilon = 0.2$  and  $\frac{\omega^2}{c_{q,EQ}^3} \kappa = 3$ .

## 5.2 Non-Linear Monte Carlo formulation

The distribution function obeys the Boltzmann equation:

$$\begin{aligned} \frac{\partial f}{\partial t} + \vec{c} \cdot \vec{\text{grad}}_{\mathcal{R}}(f) + \text{div}_{\mathcal{C}}(f \vec{a}) \\ = \int_{\mathcal{C}} d\vec{c}_* \int_{4\pi} d\vec{u} \|\vec{c} - \vec{c}_*\| \sigma(\vec{c}, \vec{c}_*) (-f f^* + f^\alpha f^\beta) \end{aligned} \quad (38)$$

where  $f \equiv f(\vec{r}, \vec{c}, t)$ ,  $f^* \equiv f(\vec{r}, \vec{c}_*, t)$ ,  $f^\alpha \equiv f(\vec{r}, \vec{c}_\alpha, t)$  and  $f^\beta \equiv f(\vec{r}, \vec{c}_\beta, t)$ .

Velocities  $\vec{c}_\alpha$  and  $\vec{c}_\beta$  are functions of  $\vec{c}$ ,  $\vec{c}_*$  and  $\vec{u}$ :  $\vec{c}_\alpha = \frac{1}{2}(\vec{c} + \vec{c}_* + \|\vec{c} - \vec{c}_*\| \vec{u})$ ,  $\vec{c}_\beta = \frac{1}{2}(\vec{c} + \vec{c}_* - \|\vec{c} - \vec{c}_*\| \vec{u})$ .

The usual PDE expression of this model, given above, can be translated into its Fredholm counterpart, following:

$$\begin{aligned} f(\vec{r}, \vec{c}, t) = \int_{-\infty}^t dt' \hat{\nu}(t') \exp\left(-\int_{t'}^t dt'' \hat{\nu}(t'')\right) \times \\ \left[ \mathcal{H}(t_I - t') f(\vec{r}_b(t_I), \vec{c}_b(t_I), t_I) \right. \\ \left. + \mathcal{H}(t' - t_I) \left(1 - \frac{\nu(t')}{\hat{\nu}(t')}\right) f(\vec{r}_b(t'), \vec{c}_b(t'), t') + \frac{s(t')}{\hat{\nu}(t')} \right] \end{aligned} \quad (39)$$

with

$$\begin{aligned} \nu(t') &= \int_{\mathcal{C}} d\vec{c}_* \int_{4\pi} d\vec{u} \|\vec{c}_b(t') - \vec{c}_*\| \sigma(\vec{c}_b(t'), \vec{c}_*) f(\vec{r}_b(t'), \vec{c}_*, t') \\ s(t') &= \int_{\mathcal{C}} d\vec{c}_* \int_{4\pi} d\vec{u} \|\vec{c}_b(t') - \vec{c}_*\| \sigma(\vec{c}_b(t'), \vec{c}_*) f(\vec{r}_b(t'), \vec{c}_\alpha(t'), t') f(\vec{r}_b(t'), \vec{c}_\beta(t'), t') \end{aligned} \quad (40)$$

where  $\vec{r}_b$  and  $\vec{c}_b$  are location and velocity corresponding to the ballistic path through  $\vec{r}$  at time  $t$  with velocity  $\vec{c}$ , such as:

$$\begin{cases} \partial_{t'} \vec{r}_b(t') = \vec{c}_b(t') \\ \partial_{t'} \vec{c}_b(t') = \vec{a}(\vec{r}_b(t')) \end{cases}$$

and  $\vec{r}_b(t) = \vec{r}$ ,  $\vec{c}_b(t) = \vec{c}$ .

The product  $f(\vec{r}_b(t'), \vec{c}_\alpha(t'), t') f(\vec{r}_b(t'), \vec{c}_\beta(t'), t')$  has been treated following the NLMC expansion exposed in Methods.

The exponential term is handled using null collisions technique, as exposed in SI4. The upper bound for cross section is set at the value it takes at the center of the gas cloud at time of maximal contraction of the undamped oscillation :

$$\hat{\nu} = \frac{\omega^3}{\left(2\pi c_{q,EQ}^2 (1 - \epsilon)\right)^{3/2}} \kappa$$

### 5.3 Algorithm

According to eq. 39, three random variables are defined for the Monte Carlo estimates:  $T'$  for  $t'$ ,  $\vec{U}$  for  $\vec{u}$  and  $\vec{C}_*$  for  $\vec{c}_*$ , with:

$$\begin{aligned} p_{T'}(t') &= \hat{\nu} \exp(-\hat{\nu}(t - t')) && \text{over } ] - \infty, t] \\ p_{\vec{U}}(\vec{u}) &= \frac{1}{4\pi} && \text{over the unit sphere} \\ p_{\vec{C}_*}(\vec{c}_*) &= p_{\mathcal{N}(\vec{u}_{Boltz}, c_{q,Boltz}^2)} && \text{over velocity space} \end{aligned} \quad (41)$$

where  $\vec{u}_{Boltz}$  et  $c_{q,Boltz}$  are respectively the mean velocity and the mean square speed at the considered time, and set to the values predicted in the case of undamped oscillation.

$f(\vec{r}, \vec{c}, t)$  can be estimated by this recursive algorithm:

**Initialisation** Sample a date  $t'$ .

**Recursion termination:** If  $t' \leq t_I$  return  $f(\vec{x}_I, \vec{c}_I, t_I)$ .

**Recursion**

Sample a velocity  $\vec{c}_*$  and a unit vector  $\vec{u}$ .

Compute the ballistic solution  $\vec{x}_b(t')$  and  $\vec{c}_b(t')$ .

Estimate  $f_* \leftarrow f(\vec{x}_b(t'), \vec{c}_*, t')$ .

Set

$$\begin{aligned} n &\leftarrow (\|\vec{c}_b(t') - \vec{c}_*\| \sigma(\vec{c}_b(t'), \vec{c}_*, \vec{u}) f_*) / (p_{\vec{C}_*}(\vec{c}_*) p_{\vec{U}}(\vec{u})) \\ Q &\leftarrow n / \hat{\nu} \end{aligned} \quad (42)$$

If  $Q \in [0, 1]$ , then set  $P \leftarrow Q$  else set  $P \leftarrow \frac{Q}{2Q-1}$ .

Sample a standard uniform  $r$ .

If  $r > P$ , then estimate  $f_b \leftarrow f(\vec{x}_b(t'), \vec{c}_b(t'), t')$  and return  $\frac{1-Q}{1-P} f_b$ ,

Else set

$$\begin{aligned} \vec{c}_\alpha(t') &\leftarrow \frac{1}{2}(\vec{c}_b(t') + \vec{c}_* + \|\vec{c}_b(t') - \vec{c}_*\| \vec{u}) \\ \vec{c}_\beta(t') &\leftarrow \frac{1}{2}(\vec{c}_b(t') + \vec{c}_* - \|\vec{c}_b(t') - \vec{c}_*\| \vec{u}) \end{aligned} \quad (43)$$

Estimate  $f_\alpha \leftarrow f(\vec{x}_b(t'), \vec{c}_\alpha(t'), t')$ .

Estimate  $f_\beta \leftarrow f(\vec{x}_b(t'), \vec{c}_\beta(t'), t')$ .

Return  $(Q/P) (f_\alpha f_\beta / f_*)$

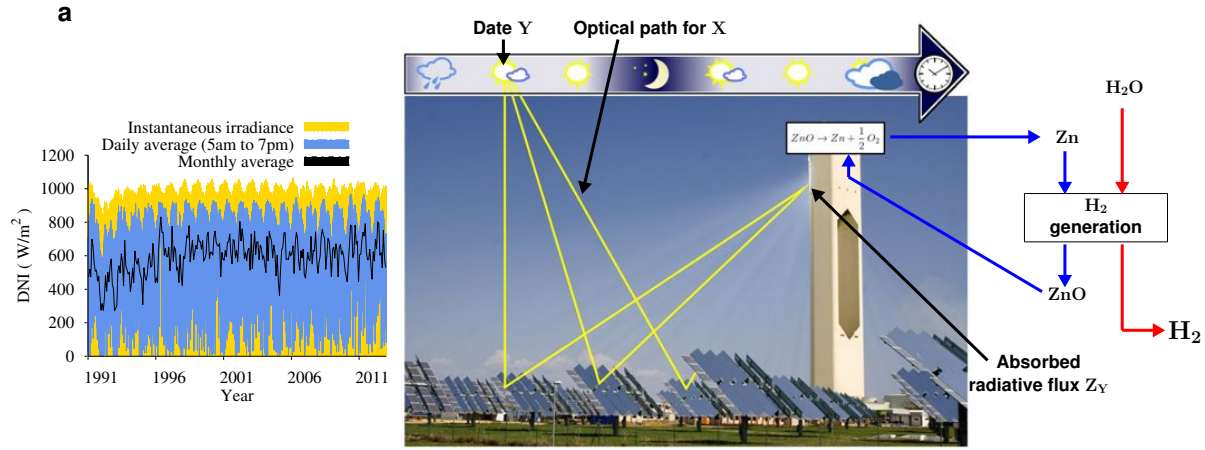
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## Extended Data Figures





**b** Time-average of chemical conversion rate at some earth location  $\mathbf{p}$

$$C(\mathbf{p}) = \mathcal{E}_Y(f(\mathbf{Z}_Y))$$

$$f(\mathbf{Z}_Y) = \exp(C_4(\mathbf{Z}_Y)^{-\alpha}) \sum_{m=0}^3 C_m(\mathbf{Z}_Y)^{(m\alpha-1)} \quad \text{Chemical conversion rate}$$

$$\mathbf{Z}_Y = \mathcal{E}_{X|Y}(\mathbf{X}|\mathbf{Y})$$

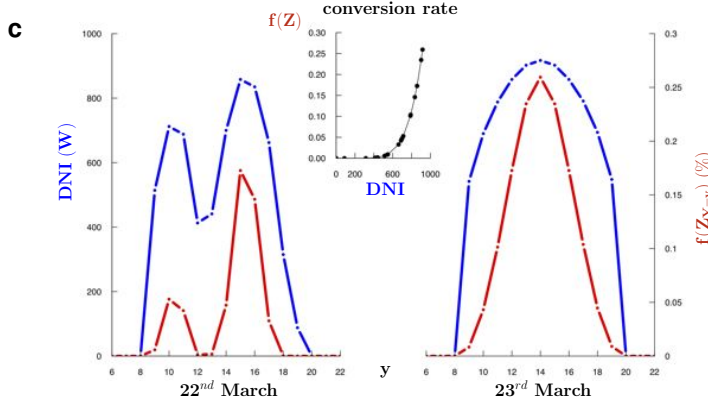
$\mathbf{Y}$

$\mathbf{X}|\mathbf{Y}$

Radiative power absorbed by the receiver

Date at the considered earth location, determining sun position and incoming solar power depending on weather conditions

Contribution at date  $\mathbf{Y}$  of sun emission along one multiple-reflection optical path leading to the receiver

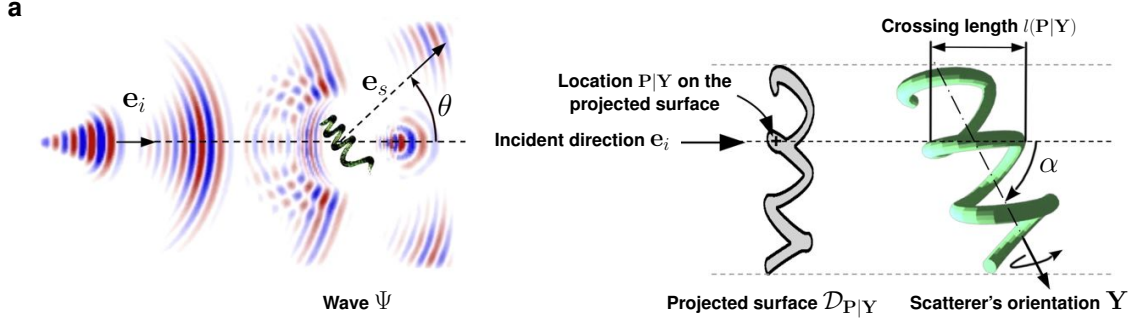


**d**

Location	$C(\mathbf{p})$
Sevilla	$0.87 \pm 0.01$
Paris	$0.85 \pm 0.01$
Albuquerque	$0.9 \pm 0.01$

**Extended Data Figure 1 | Solar thermochemical reduction of zinc oxide: conversion rate over years.** **a**, Solar-driven high temperature thermochemical cycles processes, commonly based on metal oxides reduction, are an alternative to fossil fuel-based method for  $\text{H}_2$  generation. Their practical interest depends however upon their lifetime average productivity. Here we focus on thermal reduction of zinc oxide, as the first part of a two step water splitting cycle. Photons emitted from the sun are reflected on heliostats and concentrated at the entrance of the chemical reactor in which  $\text{ZnO}$  dissociation is carried out. Solar power  $\mathbf{Z}_Y$  absorbed by the reactor at a given instant  $\mathbf{Y}$  of lifetime determines the non linear chemical conversion rate  $f(\mathbf{Z}_Y)$  of the reaction  $\text{ZnO} \rightarrow \text{Zn} + \frac{1}{2}\text{O}_2$ . **b**, Here we address the estimation of the annual solar-plant's conversion rate  $C(\mathbf{p})$  at different earth locations  $\mathbf{p}$ , by averaging the instantaneous conversion rate  $f(\mathbf{Z}_Y)$  over the statistics of

sun position and incident Direct Normal Irradiance (DNI, which fluctuates with weather). The Monte Carlo estimation combines the thermochemical knowledge of the non-linear kinetics of zinc oxide dissociation with the description of radiative transfer from the sun to a multiple-reflection central receiver solar plant. The instantaneous thermal-power  $\mathbf{Z}_Y$  collected by the receiver at moment  $\mathbf{Y}$  of the year is the average of the contribution  $\mathbf{X}|\mathbf{Y}$  of sun emissions along every optical paths. This solar power  $\mathbf{Z}_Y$  received at the entrance of the chemical reactor is used to reduce the zinc oxide with a thermochemical conversion rate  $f(\mathbf{Z}_Y)$ . **c**, Two typical days showing the received sun power (DNI, Direct Normal Irradiance) which yields  $\mathbf{Z}_Y$  after concentration by heliostats, and the conversion rate  $f(\mathbf{Z}_Y)$ ; the corresponding conversion rate from DNI to  $f(\mathbf{Z}_Y)$  is indicated in the inset. **d**, Estimated solar-plant's annual conversion rate in three earth locations. Further details are given in Supplemental Information S11.



**b** Differential scattering cross-section averaged over scatterer's orientations

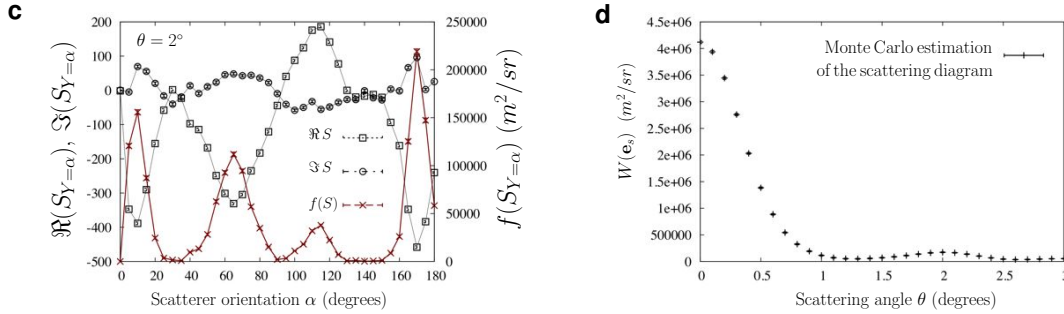
$$W(\mathbf{e}_s) = \mathcal{E}_{\mathbf{Y}}(f(S_{\mathbf{Y}}))$$

$f(S_{\mathbf{Y}}) = |S_{\mathbf{Y}}|^2$  Differential scattering cross-section

$S_{\mathbf{Y}} = \mathcal{E}_{X|\mathbf{Y}}(X|\mathbf{Y})$  Complex scattering amplitude

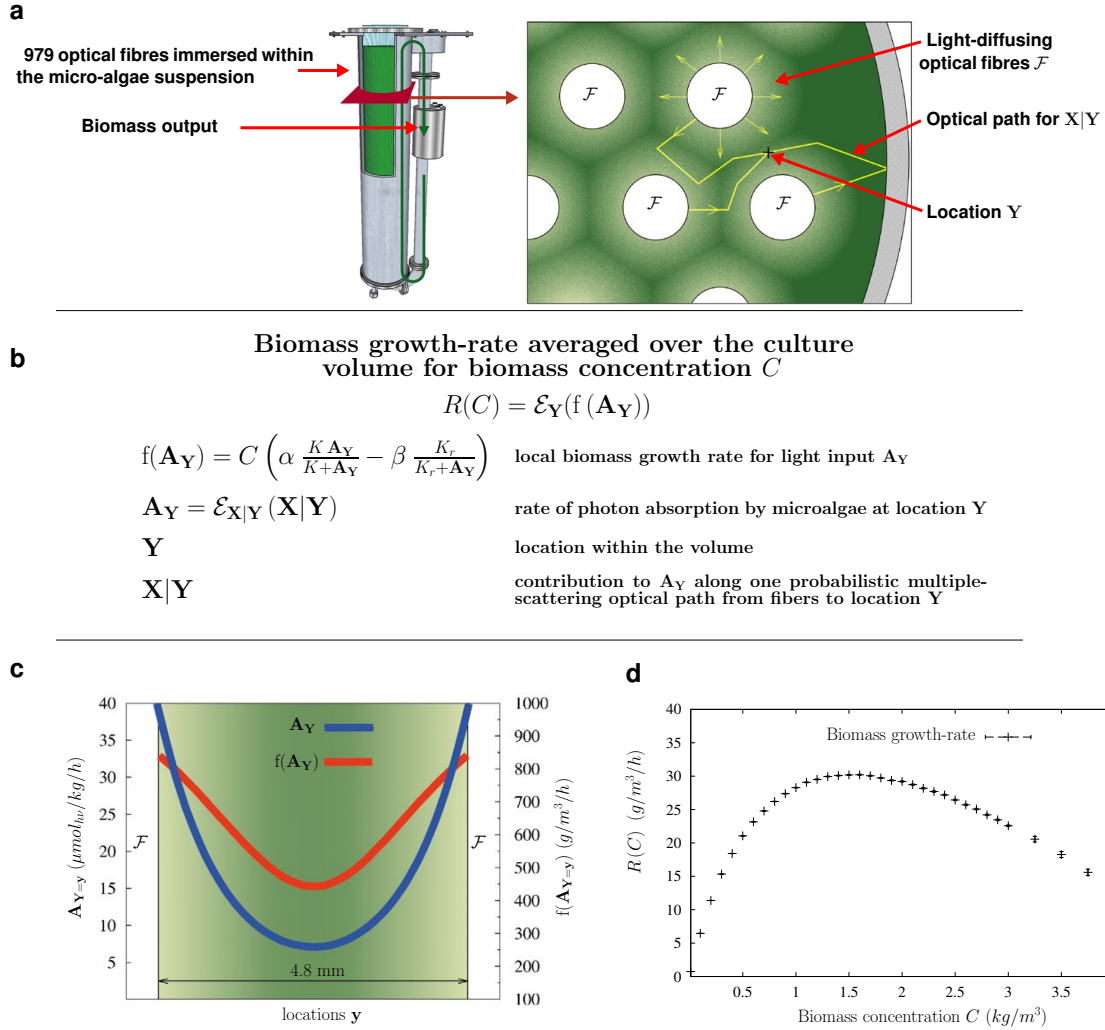
$\mathbf{Y}$  Scatterer orientation with respect to  $\mathbf{e}_i$

$X|\mathbf{Y}$  Contribution of the scattering potential through point  $\mathbf{P}$  of the projected surface



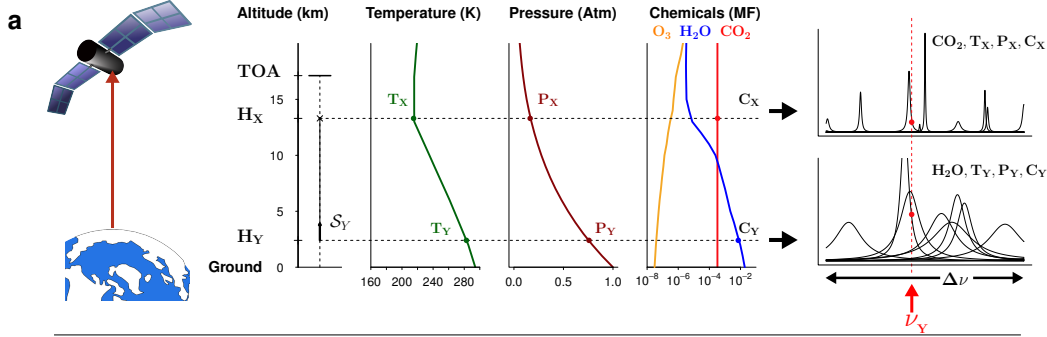
**Extended Data Figure 2 | Wave scattering by a complex-shaped and optically-soft scatterer (cyanobacterium *Arthrospira*).** **a**, An incident plane wave with propagation direction  $\mathbf{e}_i$  and wave number  $k$  is scattered by the cyanobacterium. The cyanobacterium is large compared to  $1/k$  and has low relative refractive index (optically-soft scatterer). Therefore, the scalar wave approximation is used: the field  $\Psi$  resulting from the interaction between the incident wave and the scattering potential  $U$  is solution of the scalar wave equation  $(\nabla^2 + k^2 - U)\Psi = 0$ . This wave scattering problem is identical to that of high-energy potential scattering studied by L.I. Schiff in the context of quantum mechanics. Under Schiff's approximation, the complex scattering-amplitude  $S_{\mathbf{Y}}$  (far-field region) in the forward directions depends on the scatterer orientation  $\mathbf{Y}$  through its projected surface on the plane of the incident direction  $\mathbf{e}_i$ . **b**, Here we address the estimation of  $W(\mathbf{e}_s)$  the single-scattering differential

cross-section in direction  $\mathbf{e}_s$  for a scatterer ensemble, by averaging the differential scattering cross-section  $f(S_{\mathbf{Y}}) = |S_{\mathbf{Y}}|^2$  over the statistics of orientations  $\mathbf{Y}$  (independent scattering regime). The Monte Carlo estimation of  $W(\mathbf{e}_s) = \mathcal{E}_{\mathbf{Y}}(f(S_{\mathbf{Y}}))$  at small scattering angles  $\theta$  combines the description of waves  $\Psi$  propagation with the non-linear formulation of the power that they carry  $|\Psi|^2$  (Poynting vector magnitude). The scattering amplitude  $S_{\mathbf{Y}}$  results from the interference of secondary waves contributions  $\mathbf{X}|\mathbf{Y}$  originating from the projected surface of the scatterer with orientation  $\mathbf{Y}$ . This scattering amplitude determines the differential cross-section  $f(S_{\mathbf{Y}})$  for that orientation. **c**, For a given direction  $\mathbf{e}_s$  ( $\theta = 2^\circ$ ), the scatterer orientation can greatly affect the wave shape, resulting in huge variations of the transmitted power after scattering (e.g. by five orders of magnitude from  $\alpha = 140^\circ$  to  $\alpha = 170^\circ$ ). **d**, After averaging over an isotropic orientation distribution, the cross-section depends only on  $\theta$ . Further details are given in Supplemental Information SI2.



**Extended Data Figure 3 | Phytoplankton growth in light-limited environments.** **a**, Phytoplankton is put to grow in a continuous stirred tank photobioreactor, an enclosed and perfectly controlled environment insuring optimal pH and temperature conditions, as well as non-limiting  $\text{CO}_2$  and minerals supplies to micro-algae: photosynthesis is only light-limited. Light is provided by 979 light-diffusing optical fibres  $\mathcal{F}$  immersed within the phytoplankton culture. The fibres insure a quasi-uniform light-flux density on the totality of their surface. This diluted light-input triggers an artificially sustained algal bloom with high photosynthetic efficiency and high biomass growth-rate  $R$ . The local rate of photon absorption  $\mathbf{A}_Y$  at location  $\mathbf{Y}$  determines the non-linear photosynthetic-response  $f(\mathbf{A}_Y)$  of cells at that location. **b**, Here, we address the Monte Carlo estimation of  $R(C)$  the biomass growth-rate in the culture volume as a function of biomass concentration  $C$ , by av-

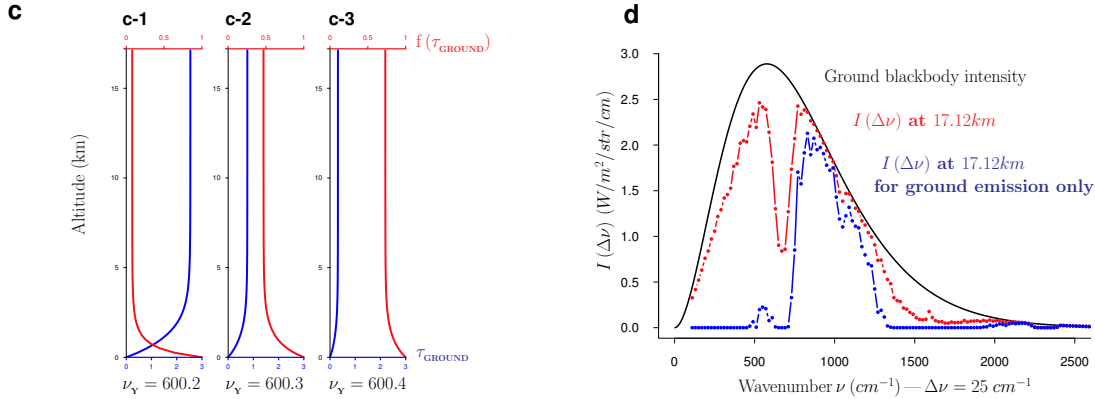
eraging the local growth-rate  $f(\mathbf{A}_Y)$  over locations in the volume. The Monte Carlo estimation of  $R(C) = \mathcal{E}_Y(f(\mathbf{A}_Y))$  combines the available knowledge of the non-linear photosynthetic growth-rate of a single cell to the description of radiative transfer within the multiple-scattering and absorbing micro-algae suspension with concentration  $C$ . The rate of photon absorption  $\mathbf{A}_Y$  at location  $\mathbf{Y}$  is the average of the contributions  $\mathbf{X}|\mathbf{Y}$  of every multiple-scattering optical paths from fibres to  $\mathbf{Y}$ . **c**, Photons absorbed by a phytoplankton-cell are non linearly converted within the photosynthetic units and the Z-scheme, leading to a spatial profile of the biomass growth-rate  $f(\mathbf{A}_Y)$ . **d**, The full-tank growth-rate is shown to depend upon the biomass concentration and indicates the optimal concentration allowing the largest biomass production rate. We note that  $R(C)$  is usually denoted by  $\langle r_x \rangle (C_x)$  in the photobioreactor literature. Further details are given in Supplemental Information SI3.



**b** Specific intensity at the Top-of-atmosphere (TOA) integrated over a spectral band  $\Delta\nu$

$$I(\Delta\nu) = \mathcal{E}_Y(\mathcal{S}_Y f(\tau_Y))$$

$\Gamma_Y$	Photon path from altitude $H_Y$ to TOA
$\mathcal{S}_Y$	Radiative source at the start of $\Gamma_Y$ for one emission-transition $L_Y(T_Y, P_Y, C_Y, \nu_Y)$ .
$f(\tau_Y) = \exp(-\tau_Y)$	Beer extinction along the path $\Gamma_Y$
$\tau_Y = \mathcal{E}_{X Y}(X Y)$	Optical thickness from altitude $H_Y$ to TOA at frequency $\nu_Y$ along $\Gamma_Y$
$X Y$	Contribution to $\tau_Y$ of one absorption-transition $L_X(T_X, P_X, C_X, \nu_Y)$ at altitude $H_X$



**Extended Data Figure 4 | Atmospheric radiative transfer: top-of-atmosphere (TOA) specific intensity (from earth toward the outer space).**

**a.** Measuring and analysing the radiation exiting an atmosphere at its top is archetypal of radiative-transfer physics in both the atmospheric science and astrophysics contexts : reflexion of solar incidence or IR-cooling to space for weather forecasting and climate-change prediction, interpretation of satellite measurements for earth monitoring, inversion of atmospheric profiles for planetary and now exo-planetary studies. For earth at infrared frequencies, photons are emitted by the ground at the surface and by the atmospheric gases at all locations  $H_Y$  between surface and top-of-atmosphere. But not all photons emitted upward reach TOA : a fraction is absorbed by the very same gases as those responsible for atmospheric emission, mainly  $CO_2$  and  $H_2O$  at all altitudes  $H_{X|Y}$  between emission-location  $H_Y$  and TOA. The emission and absorption spectra of these molecular gases display very numerous lines that are the result of energy-state transitions : observation bands of spectral width  $\Delta\nu$  typically involve several tens to hundred thousands lines of intensities and shapes that strongly depend on altitude via the atmospheric profiles of temperature  $T$ , pressure  $P$  and gaseous concentrations  $C$ .

**b.** For a given source  $\mathcal{S}_Y$  of photons travelling along a path  $\Gamma_Y$  of optical thickness  $\tau_Y$ , only  $\mathcal{S}_Y \exp(-\tau_Y)$  is transmitted and  $\tau_Y = \mathcal{E}_{X|Y}(X|Y)$  reflects the statistics of all state transitions  $X$  (absorption lines) along  $\Gamma_Y$ . But these two statistics are combined via the negative exponential function. This is the reason why standard Monte Carlo simulations of

atmospheric radiation involve a precomputation phase in which optical thickness is evaluated by adding the contributions of each absorption line in a deterministic manner: the Monte Carlo algorithm itself only deals with the statistics of photon paths. Here, on the contrary, we directly address  $\mathcal{S}_Y \exp(-\mathcal{E}_{X|Y}(X|Y))$  by non-linearly combining  $X$  and  $Y$ , i.e. both optical paths and state-transitions in one single algorithm.

**c.** Dealing with the non-linearity of Beer extinction is the leading question of band-average radiative transfer. There is a double difficulty : i) because of sharp absorption-lines, even in the narrowest bands, the optical thickness varies of orders of magnitude with frequency (compare  $\tau_Y$  at three frequencies, c-1, c-2, c-3); ii) because of the variations of pressure and composition with altitude, the spatial dependance of  $\tau_Y$  is difficult to handle, and so is the diversity of its exponential translation :  $f(\tau_Y) \approx 1 - \tau_Y$  in c-3 whereas it is highly non-linear c-1.

**d.** When applied to a typical state of earth atmosphere, our NLMC algorithm addresses successfully the simulation of band-averaged outgoing IR-radiation (zenith angle,  $\Delta\nu = 25 cm^{-1}$ ) without the need of pre-processing the state-transition database (here Hitran) : for each sampled path, several state-transitions are sampled in the database, depending on the sampled order of the Taylor expansion of the exponential (using the null-collision approach). In red, the band-averaged specific intensity for several bands covering the whole infrared ; in blue, the part of this intensity due to photons emitted at the surface (close to red at frequencies where the atmosphere is nearly transparent). Further details are given in Supplemental Information SI4.